

Embedding sea surface temperature anomalies into the stock recruitment relationship of red mullet (*Mullus barbatus* L. 1758) in the Strait of Sicily*

D. LEVI, M.G. ANDREOLI, A. BONANNO, F. FIORENTINO, G. GAROFALO,
S. MAZZOLA, G. NORRITO, B. PATTI, G. PERNICE, S. RAGONESE,
G.B. GIUSTO and P. RIZZO

Istituto di ricerche sulle Risorse Marine a l' Ambiente (IRMA-CNR), Via Luigi Vaccara 61, 91026 Mazara del Vallo (TP), Italy. E-mail: levi@irma.pa.cnr.it

SUMMARY: In the Strait of Sicily, red mullet (*Mullus barbatus* L. 1758) is one of the most important resources for the bottom trawlers exploiting the shelf fishing grounds. At present, the younger age groups (groups 0-1) represent most of the catches. In order to give more effective advice for fisheries management in the area, this paper is aimed at further investigating the stock-recruitment relationship for red mullet by including environmental information. Indices of the mean abundance of red mullet older than 1 year (spawners) and younger (recruits) were obtained from autumn trawl survey samples by analyzing the length-frequency distributions. Sea surface temperature (SST), expressed as the average monthly anomalies over the period from May to September (referring to 1961-90 climatology), was considered as a proxy for oceanographic processes affecting recruitment. An exploratory analysis was performed using three stock-recruitment relationships (Cushing, Ricker and Beverton-Holt), with and without the effect of SST. The best fit ($R^2=0.80$) was obtained with a Ricker modified curve ($\ln R = \ln(70.17113) + \ln S - 0.063812S + 0.060544(S \cdot E)$), including SST anomalies in July and August as a masking factor. Results showed that, for a given spawning stock level, higher recruitment levels corresponded to warmer than average SST in early life stages.

Key words: red mullet, *Mullus barbatus*, stock-recruitment relationships, sea surface temperature, Strait of Sicily, Mediterranean.

INTRODUCTION

It is widely recognized that Mediterranean demersal fisheries exploiting several species use various types of gear, but mainly prefer small mesh size trawling (Caddy, 1990; Farrugio and Papaconstantinou, 1998). The majority of the catch includes fish in the first two years of life, with high exploitation rates even on the 0 group (Caddy, 1998).

Since many Mediterranean demersal species of fish and cephalopods are recruited to fishery as 0

group, understanding the relationships between recruitment strength, spawning stock size and environmental factors appears crucial in order to provide short-term forecasts so that managers can adjust fishing regulations in response to changes in stock size.

It is well known that spawner stock and recruitment relationships (SRR) are one of the most complex issues in the dynamics of exploited resources. In fact, numerous factors may confound these relationships because of the nature of the phenomena and due to bias in data collection (Sissenwine and Shepherd, 1987; Hilborn and Walters, 1992).

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The role of environmental variability in recruitment dynamics of exploited populations has long been recognized and many investigations on this issue can be found in the literature, e.g. Cushing (1982), Glantz (1992), Laevastu (1993), Chambers and Trippel (1997). However, the association between environmental factors and recruitment variability has often been questioned: because of the large number of environmental variables that may be tested for correlation with recruitment, the risk of a significant spurious correlation is high (Botsford, 1987; Walters and Collie, 1988; Hilborn and Walters, 1992).

The aim of this study was to study a specific case of SRR in the Mediterranean, incorporating information on environmental variables affecting recruitment in the model, as suggested by Iles and Beverton (1998). This procedure should enable a simultaneous analysis of the effect of spawners and environmental factors on recruitment, reducing the unexplained variability and improving our understanding of recruitment dynamics.

In this context, the Strait of Sicily represents an excellent test site for studying coupling phenomena between the main oceanographic processes and biological resources because of its limited geographic scale.

Because of its widely-recognized importance and the relative ease with which data can be obtained, sea surface temperature (SST) was considered as an environmental factor potentially capable of affecting inter-annual fluctuations in fish abundance (Templeman and Fleming, 1953; Sissenwine, 1984; Garrod and Shumacher, 1994; Southward and Boalch, 1994). The effect of SST on recruitment is well-established for some temperate fish stocks (e.g. Northeast Arctic Cod), but the processes involved are complex (Brander, 1997).

SST could directly affect the early stages of life cycles in fish populations (Brett, 1970; Boeuf, 1988; Wood and McDonald, 1997), but variations in this parameter may also reflect changes in oceanographic processes, e.g. current transport and turbulent mixing, affecting food and larval retention processes. More and more papers suggest that the latter are important in regulating the strength of recruitment (Nelson *et al.*, 1977; Parrish *et al.*, 1981; Lasker, 1981; Bailey, 1981; Bakun, 1985; Corten, 1986; Buckley and Lough, 1987; Peterman and Bradford, 1987; Kennet *et al.*, 1989; Pihl, 1990; Van der Veer *et al.*, 1998; Daskalov, 1999; Sanchez and Gil, 2000).

The biological data presented in this study concern red mullet, *Mullus barbatus* (L., 1758), one of the most important resources for shelf demersal fishery (down to 200m) in the Mediterranean (Levi *et al.*, 1993; Fiorentini *et al.*, 1997). The investigation was restricted to the red mullet inhabiting the Sicilian side of the Strait, because the biological features of those living on the Tunisian shelf suggest that they belong to another population (Levi *et al.*, 1992; 1994).

The biological characteristics (sexual maturity at 1 year old) and exploitation features (similar vulnerability to trawling from recruitment onward) make red mullet an interesting case for studying the renewal capability of resources in the Mediterranean.

The time series of abundance indices for spawners and recruits used in this paper were obtained as part of programs for evaluating demersal resources in the Italian seas from 1985 onwards (Relini, 1998).

ENVIRONMENTAL AND BIOLOGICAL BACKGROUND

The Strait of Sicily connects the two main, western and eastern basins of the Mediterranean. At its narrowest section, between Cape Bon (Tunisia) and Mazara del Vallo (Sicily), it is about 130 km wide (Fig. 1).

Like the Strait of Gibraltar, it is characterized by a two-layer flux model. The upper layer, called "Modified Atlantic Water" (MAW), is identified by water with a relatively low salinity flowing from the western to the eastern basin. The lower layer, called "Levantine Intermediate Water" (LIW), is identified by water with a relatively high salinity flowing in the opposite direction, like an undercurrent (Bethoux, 1980; Manzella *et al.*, 1988; Onken and Sellshopp, 1998).

The complex bathymetry influences the features of the currents in the region. A crest in the middle of the strait divides the flow into two parts: one along the Sicilian shelf, which is narrow and deep, the other off the Tunisian coast, which is wide and shallower. Astraldi *et al.* (1996) showed that these two parts have two different flows, both for MAW and for LIW. Measurements have shown that the mesoscale signal prevails in the upper layer, with the presence of eddies and meanders.

According to Robinson *et al.* (1991) the MAW motion, called the "Atlantic-Ionian Stream" (AIS),

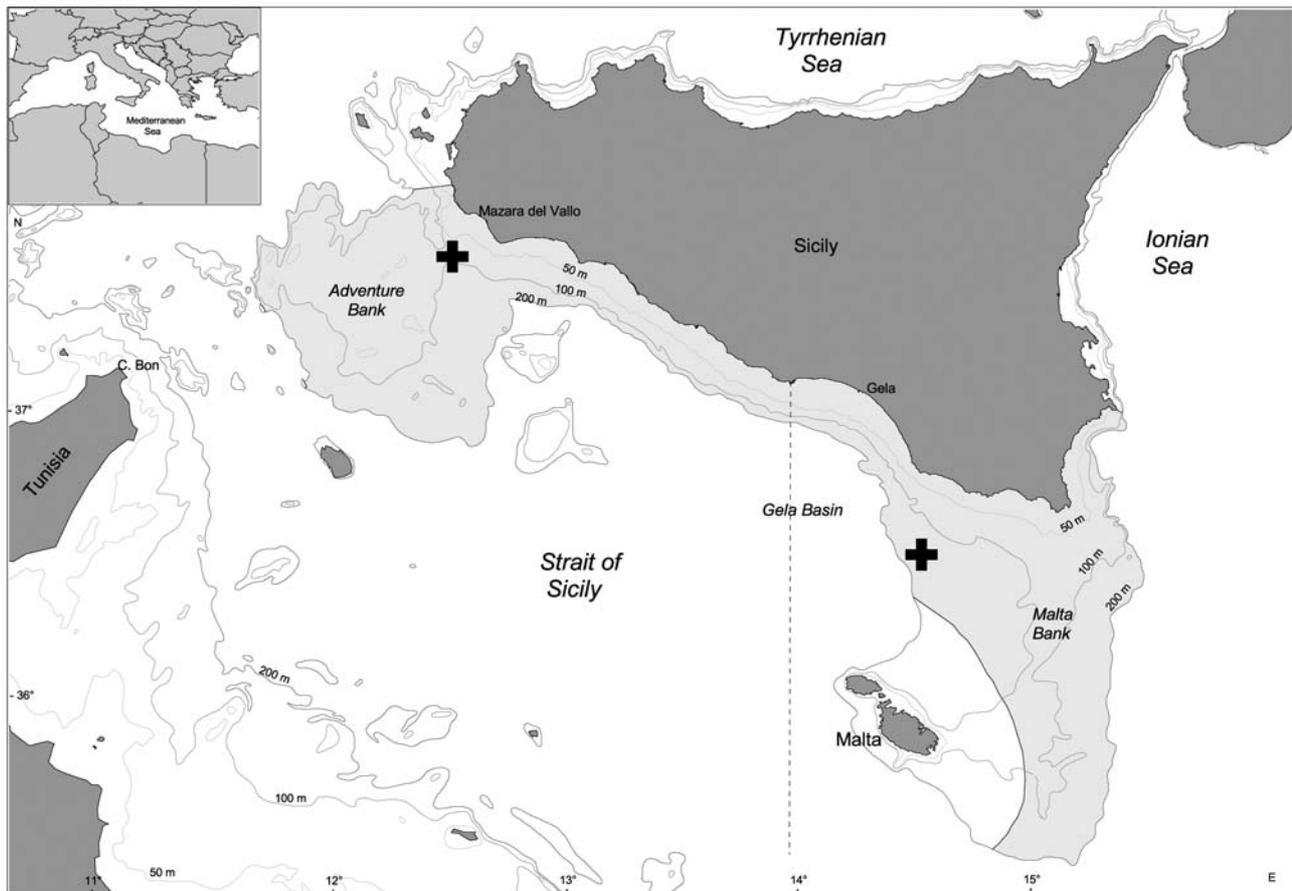


FIG. 1. – The study area with the main bathymetry and the line separating the two sectors. Black crosses represent the position of oceanographic stations.

has quite a steady mean path. It enters the strait from the western boundary along the Adventure Bank (south-western shelf of Sicily), coming close to the Sicilian shore in the middle of the south coast of Sicily (the deep Gela basin) and separates again when it encounters the Malta Bank (south-eastern shelf of Sicily).

The AIS encircles two large cyclonic meanders, identifying two sectors, one around the Adventure Bank on the western side, the other around the Malta Bank to the east. This is illustrated in Figure 2, taken from Mazzola *et al.* (2000).

A preliminary analysis of the SST at two sites roughly representative of the two sectors was carried out on the basis of 1961-90 climatology data. In particular, the mean annual SST estimated for the western and eastern sectors amounted to 19.3°C ($\pm 4.2^\circ\text{C}$) and 19.9°C ($\pm 4.1^\circ\text{C}$), respectively.

This kind of water circulation favors permanent upwelling to the left of the AIS, which is reinforced by westerly winds (Piccioni *et al.*, 1988; Mazzola *et al.*, 1998).

The offshore transport associated with this upwelling may affect the survival/mortality rates in the early life stages of red mullet, influencing the strength of recruitment.

Since eggs, larvae and post-larvae up to 30-35 mm of *M. barbatus* are pelagic and live in the surface waters (Lo Bianco, 1908-1909; Montalenti, 1933), the circulation of MAW, with its meanders and upwellings, is assumed to affect red mullet recruitment.

MATERIAL AND METHODS

SST data

Monthly 1°x1° (spatial resolution) global SST analyses, available from November 1981 to the present day, were obtained from a collection of SST analyses prepared at the National Center for Environmental Prediction (formerly NMC) by D. Reynolds, D. Stokes, and T. Smith. These analyses

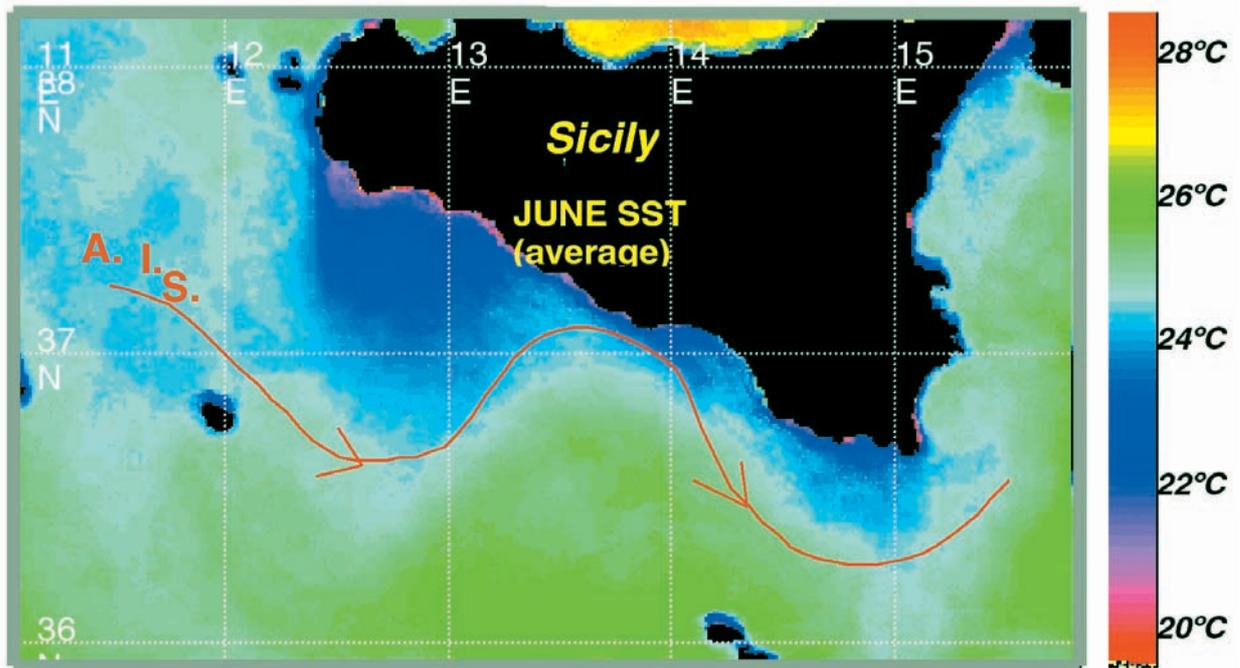


FIG. 2. – Sea Surface Temperature map from NOAA-14 satellite in June 1998. It is a monthly-averaged situation and must be interpreted in terms of the steadiness of certain mesoscale features. The presence of the large blue regions (cool surface water) off Mazara del Vallo (south-western region of Sicily) and off Capo Passero (southernmost tip of the island) indicate fairly permanent upwelling areas that fit well with descriptions of the general circulation within the Strait. The red line indicates an “averaged” path of the Atlantic-Ionian Stream (AIS) (from Mazzola *et al.*, 2000).

were determined by blending marine surface observations and satellite AVHRR data using an optimum interpolation (OI) method. A description of the OI analysis can be found in Reynolds and Smith (1994). SST data were provided by the Data Support Section of the University Corporation for Atmospheric Research (UCAR), Boulder, Colorado, USA, from their web site at <http://dss.ucar.edu/>.

From the whole data set, we extracted two SST time series relating to the sites off the southern coast of Sicily shown in Figure 1 (geographical coordinates: 37°30'N-12°30'E and 36°30'N-14°30'E), located roughly in the middle of the western and eastern sectors of the study area, where red mullet recruitment is seen to occur.

The Smith and Reynolds (1998) adjusted OI monthly climatology data (base period 1961-1990) for the two above-mentioned locations were used to generate both average monthly SST and monthly anomalies. We restricted our analysis to monthly SST anomalies from May to September because peak spawning for red mullet occurs in May (Levi, 1991) and our aim was to test the influence of the environmental factor (SST anomalies) on the early life stages of the species, before its subsequent recruitment to the fishery.

Stock-Recruitment data

Samples of red mullet were collected as part of the Italian Programs for evaluating demersal resources in the Strait of Sicily (Levi *et al.*, 1998). Bottom trawl surveys were carried out on an area of about 50,000 km², between 10 and 800 m in depth, using a random stratified design. The number of hauls per stratum was proportional to the area of each stratum. The period under study was 1985-98, during which, among others, 6 spring and 9 autumn surveys were carried out.

On the basis of information on spawning and recruitment periods in the area (Levi, 1991), the fish were divided into two groups, i.e. the young of the year forming the 0 group (born in spring and recruited in summer) and the adults (fish older than 1 year), which spawned in spring.

It was assumed that the recruits found on the bottom along southern Sicilian coasts during the autumn of a given year derive from the adults that spawned during the spring of the same year. Since the spring has not been sampled regularly, the spawner indices used in the calculation were also derived from autumn surveys, after verifying that the abundance of spawners after 6 months was proportional to their abundance at spawning time.

Given the two different meanders around the Adventure and Malta banks, characterized by different average SST, the length-frequency distributions (LFD) were analyzed separately for the western and eastern hauls, dividing the Strait of Sicily as shown in Figure 1.

The identification and strength of age groups in the length-frequency distributions were obtained in 3 steps.

First the overall LFD for each survey was analyzed in order to split the first component and identify the mean length (ML) and standard deviation (SD) of recruits for the year. The LFD components were separated using the NormSep maximum likelihood procedure, as implemented in the FiSAT package (Gayanilo *et al.*, 1996). The preliminary seed values required were obtained by applying the Bhattacharya procedure, again in FiSAT. Then the number of recruits for the year per km² (R) was calculated for each haul as individuals having a total length from the smallest up to ML + 1SD. Finally, the corresponding biomass of adults per km² (S) was simply computed by summing the individual weights of adults.

The simple mean indices of recruits and spawners per km² for each survey were used directly for modeling purposes.

In order to show the highest-density areas of both recruits and spawners, data were pooled over the years and relative densities were calculated as ratios with

respect to the maximum for each year and sector.

All maps were produced with the ArcView-GIS software by ESRI.

Analytical techniques

Three different classes of stock–recruitment relationship were tested, namely the two-parameter models of Cushing, Ricker and Beverton-Holt, aiming for the best fit of the available data, both with and without incorporating the SST information in the equations. The models reflect the stock-recruitment relationships reported in Iles and Beverton (1998) and are all listed in Table 1. The classification of the environmental influences, given in the table, is based on the work by Neill *et al.* (1994), who state that environmental factors affecting recruitment can be divided into five kinds: controlling, limiting, lethal, masking and directive. The effects of these factors are broadly discussed by Iles and Beverton (1998). The root conceptual framework can be found in Fry (1947).

In the present study, only three kinds of effect were considered, i.e. controlling, limiting and masking. The first is caused by all factors that alter the rate of change in the numbers of young fish in time; the second by factors that vary the carrying capacity of the habitat for recruits; the third is due to factors that modify the rate of recruitment for any given spawner stock size.

TABLE 1. – Type of S-R models with and without environmental factors affecting recruitment according to Iles and Beverton (1998). Adjusted determination coefficients (R²) and estimated equations obtained for different models both including and excluding SST effects (average value of SST anomalies over July and August). Equation (6) shows the best fit to the data, accounting for about 80% of the total variance. For the purposes of comparison, the models listed in this table have all been fitted in the ln domain by non-linear regression. R² values of models having at least one parameter not significantly different from zero are in italics.

Type of SRR	Equation		Inclusion of the environmental factor	Implicit role of the environmental factor	Estimated equation	R ²
Cushing	$R = \alpha S^{\gamma}$	(1)	no	-	$\ln R = \ln(68.28998) + 0.875666 \ln S$	<i>0.57</i>
Cushing modified	$R = \alpha S^{\gamma} e^{cE}$	(2)	yes	Controlling effects	$\ln R = \ln(61.50826) + 0.776993 \ln S + 0.607726E$	<i>0.70</i>
Cushing modified	$R = \alpha S^{\gamma+cE}$	(3)	yes	Masking effects	$\ln R = \ln(79.55701) + 0.633832 \ln S + 0.310737E \cdot \ln S$	<i>0.74</i>
Ricker	$R = \alpha S e^{-\beta S}$	(4)	no	-	$\ln R = \ln(67.68564) + \ln S - 0.02645^{(*)} S$	<i>0.60</i>
Ricker modified	$R = \alpha S e^{-\beta S} e^{cE}$	(5)	yes	Controlling effects	$\ln R = \ln(54.07882) + \ln S - 0.035348^{(**)} S + 0.618332E$	<i>0.74</i>
Ricker modified	$R = \alpha S e^{-\beta S(1+cE)}$	(6)	yes	Masking effects	$\ln R = \ln(70.17113) + \ln S - 0.063812S + 0.060544(S \cdot E)$	<i>0.80</i>
Beverton-Holt	$R = S/(b + aS)$	(7)	no	-	$\ln R = \ln(S/[0.015002 + 0.000434^{(***)} S])$	<i>0.59</i>
B-H modified	$R = S e^{cE} / [b + aS]$	(8)	yes	Controlling effects	$\ln R = \ln(S/[0.018309 + 0.000831^{(*)} S]) + 0.612632E$	<i>0.73</i>
B-H modified	$R = S / [b + a e^{cE} S]$	(9)	yes	Limiting effects	$\ln R = \ln(S/[0.01283 + 0.001445 e^{-1.76158^{(**)} E} S])$	<i>0.74</i>
B-H modified	$R = S / [b e^{cE} + aS]$	(10)	yes	Masking effects	$\ln R = \ln(S/[0.02078 e^{-0.63956^{(**)} E} + 0.000384^{(***)} S])$	<i>0.70</i>

R: recruits, S: spawners, E: SST

(*) parameter not significantly different from zero ($0.1 < p < 0.3$).

(**) parameter not significantly different from zero ($0.05 < p < 0.1$).

(***) parameter not significantly different from zero ($0.3 < p < 0.4$).

For the purposes of comparison, the models considered have all been fitted in the ln domain by non-linear regression.

As for the models incorporating the environmental variable, various combinations of the available monthly SST anomalies were tested to find the best data fit (in terms of higher proportions of variance accounted for), starting from the information for single months and computing all possible average values for 2, 3, 4 and 5 consecutive months, i.e. the average SST anomaly was calculated for: May-June

(MJ), June-July (JJ), July-August (JA), August-September (AS), May-July (MJJ), June-August (JJA), July-September (JAS), May-August (MJJA), June-September (JJAS), and May-September (MJJAS). The selected variable was consequently included in all models incorporating an environmental effect. After fitting the full models and estimating the parameters, the statistical hypothesis that all these parameters differed significantly from zero was tested.

All statistical analyses were carried out using the STATISTICA for Windows software package.

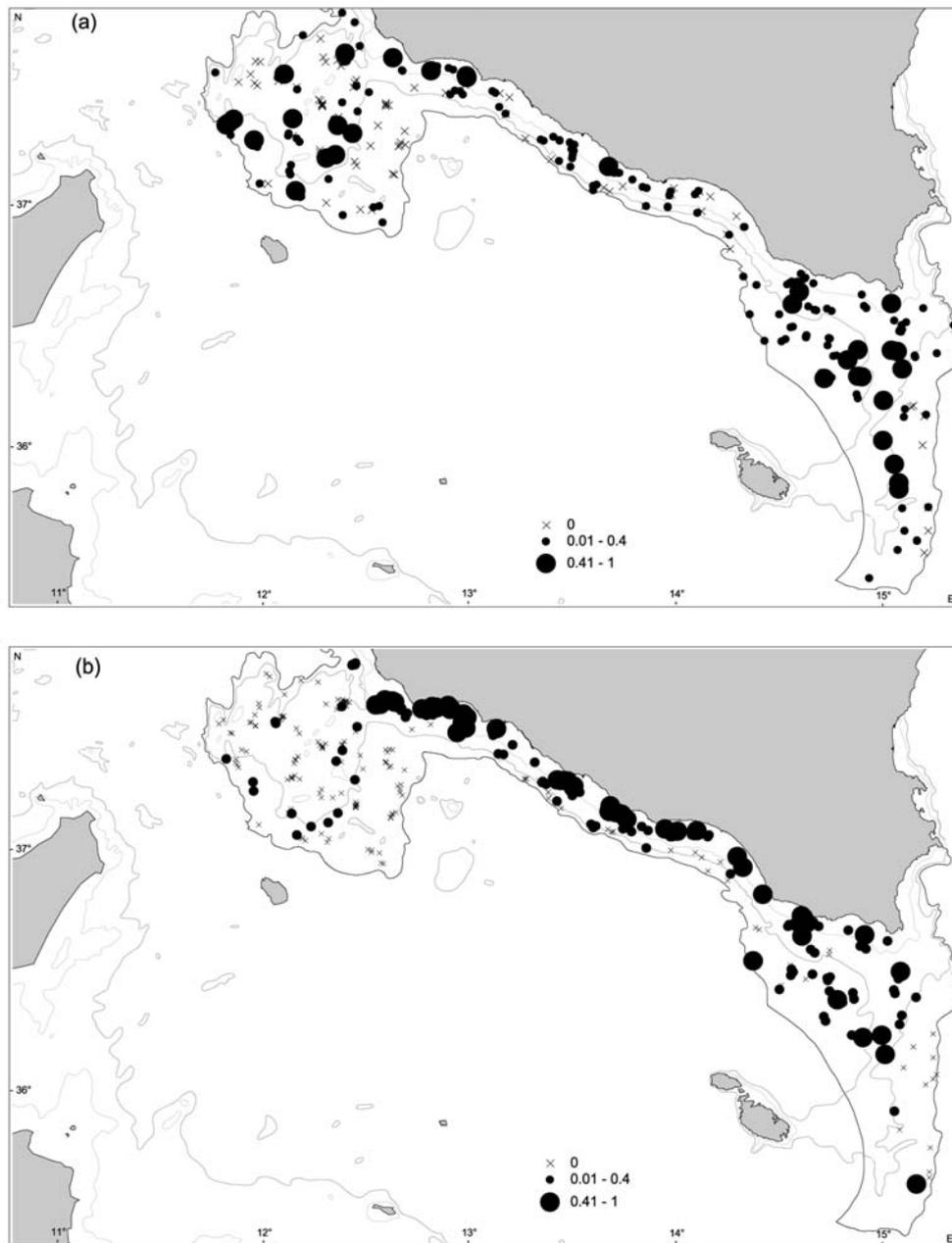


FIG. 3. – Maps of the relative densities of spawners from spring surveys (a) and recruits from autumn surveys (b). In order to show the highest-density areas of both recruits and spawners, data were pooled over the years and relative densities were calculated as ratios with respect to the maximum for each year and sector.

RESULTS

The analysis of SST data revealed different temperature patterns for the two sectors of the study area, the SST in the west being consistently lower than in the east. The mean monthly difference was -0.557, a value which departs significantly from 0, as demonstrated by applying Student's paired t-test ($t = -15.139$, $df = 11$, $p > 0.005$).

Further support for the W-E differentiation of the study area emerged from the synthetic representation of the main nursery and spawning areas, given in Figure 3. The maps show the relative densities of both spring spawners and autumn recruits, and reveal two distinct spawning areas roughly corresponding to Adventure Bank (western sector) and Malta Bank (eastern sector). Spring spawners were more abundant at deeper levels (down to 200 m), though in the western sector they even occurred within 50 m depth (Fig. 3a). Conversely, recruits from autumn surveys were more abundant on very shallow bottoms, but they may be found even to around 100 m in depth in the eastern sector (Fig. 3b).

The abundance of adults in autumn correlated significantly with the one found in spring ($r = 0.56$;

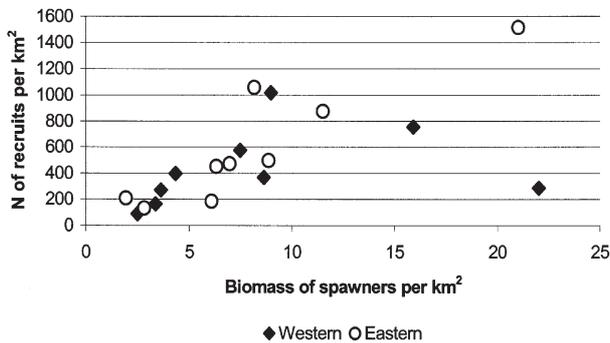


FIG. 4. – Stock-recruitment scatter plot.

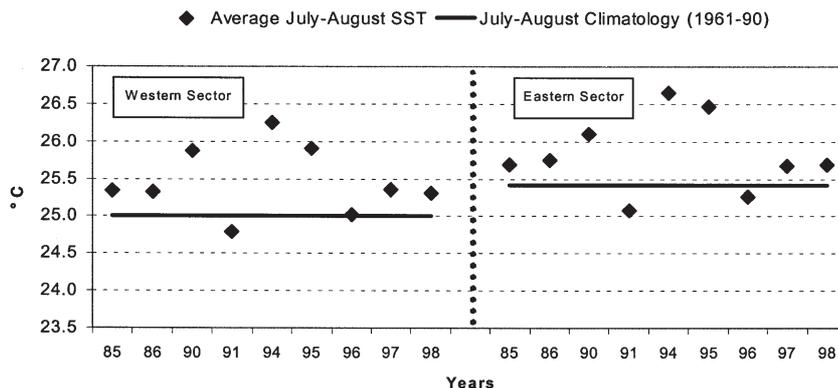


FIG. 5. – Average July-August SST time series and corresponding climatology (base period: 1961-90) relating to the two sites off the southern coast of Sicily shown in Fig. 1, approximately representative of the western and eastern sectors of the study area. The SST anomaly series for each sector were calculated as differences between the plotted variables.

TABLE 2. – Parameters and standard errors of the estimated Ricker model (Eq. 6 in Table 1) incorporating the environmental factor as a masking effect.

	α	β	c
Model: $\ln(R) = \ln(\alpha) + \ln(S) - \beta \cdot S + c \cdot S \cdot \text{SSTanom}$			
Dep. var: $\ln(R)$ Loss: $(\text{Obs-Pred})^2$			
Final loss: 1.807 R=0.89 Variance explained: 80%			
Estimate	70.17113	0.063812	0.060544
Std.Err.	12.99323	0.020141	0.017591
t(12)	5.400593	3.168249	3.44172
p-level	0.00016	0.008096	0.004879

$p < 0.05$). These findings enabled the autumn abundance indices to be used for S-R studies.

The scattering of recruitment in relation to stock size, illustrated in Figure 4, shows a fairly clear link between the variables, though the relationship tends to weaken at higher stock levels.

As a result of the test to detect the best combination of monthly SST anomaly information to include in stock-recruitment relationships, the average value over the July-August period was selected. This is shown in Figure 5 for the period under study (1985-1998). It is worth noting that, during this period, the average SST were consistently lower in the western sector than in the east, but generally higher than the corresponding climatological values.

Table 1 lists the estimated equations for all the models, limited to those including the selected variable. Equation (6) emerges as the best fit to the data, accounting for about 80% of the total variance. Results suggest that introducing SST information enables us to justify up to 20% more of the variance than the 60% maximum obtained when the environmental factor is not considered. It is also worth noting that the estimation of the β parameter of the Ricker model (Eq. 4) did not differ significantly from zero, whereas in the modified Ricker model (Eq. 6) all the parameters were highly significant ($p < 0.01$, Table 2).

$$\text{Model: } \ln(R) = \ln(\alpha) + \ln(S) - \beta \cdot S + c \cdot S \cdot \text{SSTanom}$$

$$\ln(R) = \ln(70.17113) + \ln(S) - (0.0638123) \cdot S + (0.0605436) \cdot S \cdot \text{SSTanom}$$

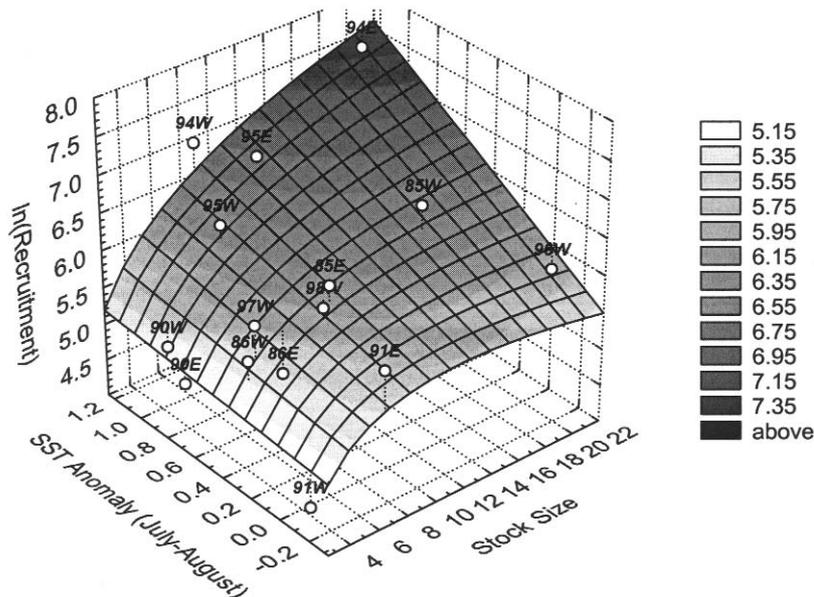


FIG. 6. – Recruitment (number of fish per km²) as a function of stock size (kg per km²) and average July-August SST anomaly (°C). SRR according to the Ricker model including SST anomalies as a masking effect (Eq. 6 in Table 1).

Figure 6 shows that the SST anomalies act as a moderating factor capable of depressing or boosting the recruitment corresponding to a given stock size level resulting from the application of a simple Ricker model alone. In particular, the estimated recruitment appears to be proportional to the SST anomaly, but the influence of the environmental factor grows with increasing stock size levels.

DISCUSSION

All the literature on environmental effects on recruitment, including the papers quoted in the introduction, refers to Atlantic stocks, mainly gadoids (Sanchez and Gil, 2000 and literature therein; Brander, 2000 and literature therein; papers in this volume), or to small pelagics.

Historically, maybe, the focus on both groups was mainly generated by the inability of “traditional” stock assessments to prevent severe collapses.

On the other hand, monitoring and possibly forecasting recruitment is of paramount importance in the Mediterranean, since a large proportion of catches consists of age groups 0 and 1. This is the first time that environmental factors have been incorporated in stock-recruitment models to identify more realistic relationships for Mediterranean resources.

Results show that investigating SRR, taking the simultaneous effects of spawning stock and SST into account, may be crucial to short-term forecasting. Although the present configuration of the model (using autumn adult abundance) is inadequate for predicting the strength of recruitment, monitoring SST in July and August can be used to regulate fishing effort on the 0 group in the autumn in order to prevent recruitment over-fishing risks, taking the likely exploitation of the resources into account (Levi *et al.*, 1993).

The first evidence of the dependence of red mullet recruitment on spawner abundance in the Mediterranean came from Fiorentino *et al.* (1998) and Zamboni *et al.* (2000) for the Ligurian Sea (north western Mediterranean). They modeled relationships with the classical Ricker and Beverton and Holt models without considering any effects of environmental parameters, showing a great variability in the scattering of recruitment around the regression model, possibly due to errors in measuring stock and recruitment and to variations in environmental factors.

As regards the red mullet of the Strait of Sicily, the Ricker model with SST anomalies affecting recruitment as a “masking factor” (for July-August) exhibited the best fit to data ($R^2=0.80$).

Results suggest that positive SST anomalies positively affect recruitment.

Available data do not support the claim that SST modulates recruitment strength by affecting the development of early juvenile stages (direct or physiological control) or regulating their feeding and settlement (indirect or ecological control).

However, since spawning in the area occurs mainly in May-June at depths between 50 and 200 m (Fig. 3a) and recruitment occurs on bottoms very close to the coast in August-September (10-30m, Fig. 3b), a decreased offshore transport can be postulated as a factor favoring the success of recruitment. Assuming that SST directly modulates the survival rates of pelagic eggs and larvae (Ellertsen *et al.*, 1989; Pepin, 1991), SST anomalies in May-June might be expected to be more important in SRRs than in July-August. Given that recruitment dynamics were positively influenced by July-August high anomalies (warmer than average surface water), it seems more reasonable to consider a reduced upwelling regime and a consequent lower level of offshore transport (testified by high SST anomaly values) as the key factor capable of supporting the survival of juvenile stages during their migration towards the coastal nursery areas. Conversely, lower than average SST values could be a proxy of processes capable of enhancing pre-recruit mortality rates directly (offshore transport obstructing migration) and/or indirectly (poor food availability induced by turbulent mixing of sea surface layers).

Future research will concentrate along two main lines.

On the one hand, more data (e.g. satellite as in Demarcq, 2000) and longer time series are needed before drawing firmer conclusions and further reducing uncertainty. These results must be considered with caution because the good fit to 1985-1998 data, obtained by including environmental parameters in the SRR, has to be validated in future investigations. The relationships between SST, upwelling and offshore currents in the area also need to be better defined and/or maybe synthesized (Bakun, 1996).

On the other hand, while fishing seemed to suggest the existence of two separate spawning areas, one for the Adventure Bank and the other for the Malta shelf, such evidence was merely used here to strengthen a general model for red mullet on the shelf of southern Sicily. Whether two meta-populations inhabit the two aforesaid areas should also be clarified by further research.

It is also to be hoped that parallel work, e.g. on small pelagics and the environment (Mazzola *et al.*,

1998, 2000; Garcia Lafuente *et al.*, 2001) in the same areas will improve forecasting by explaining factors involved in the causal chain (plankton biomass and seston concentrations, permanence of local gyres and meanders) on the mesoscale.

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